Baseline Prolonged PR Interval and Outcome of Cardiac Resynchronization Therapy: A Systematic Review and Meta-analysis

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Abstract

Background: Recent studies suggest that baseline prolonged PR interval is associated with worse outcome in cardiac resynchronization therapy (CRT). However, a systematic review and meta-analysis of the literature have not been made.

Objective: To assess the association between baseline prolonged PR interval and adverse outcomes of CRT by a systematic review of the literature and a meta-analysis.

Methods: We comprehensively searched the databases of MEDLINE and EMBASE from inception to March 2017. The included studies were published prospective or retrospective cohort studies that compared all-cause mortality, HF hospitalization, and composite outcome of CRT with baseline prolonged PR (> 200 msec) versus normal PR interval. Data from each study were combined using the random-effects, generic inverse variance method of DerSimonian and Laird to calculate the risk ratios and 95% confidence intervals.

Results: Six studies from January 1991 to May 2017 were included in this meta-analysis. All-cause mortality rate is available in four studies involving 17,432 normal PR and 4,278 prolonged PR. Heart failure hospitalization is available in two studies involving 16,152 normal PR and 3,031 prolonged PR. Composite outcome is available in four studies involving 17,001 normal PR and 3,866 prolonged PR. Prolonged PR interval was associated with increased risk of all-cause mortality (pooled risk ratio = 1.34, 95% confidence interval: 1.08-1.67, p < 0.01, I² = 57.0%), heart failure hospitalization (pooled risk ratio = 1.30, 95% confidence interval: 1.16-1.45, p < 0.01, I² = 6.6%) and composite outcome (pooled risk ratio = 1.21, 95% confidence interval: 1.13-1.30, p < 0.01, I² = 0%).

Conclusions: Our systematic review and meta-analysis support the hypothesis that baseline prolonged PR interval is a predictor of all-cause mortality, heart failure hospitalization, and composite outcome in CRT patients. (Arq Bras Cardiol. 2018; [online].ahead print, PP.0-0)

Keywords: Heart Failure/complications; Heart Conduction System/physiopathology; Ventricular Dysfunction/complications; Cardiac Resynchronization/methods; Review; Meta-Analysis.

Introduction

It has been widely accepted that surface electrocardiogram findings are associated with prognosis in congestive heart failure (HF) patients who have required cardiac resynchronization therapy (CRT), particularly the QRS complex. QRS duration and morphology is a well-established predictor of outcome among patients receiving CRT as well as selection criteria for CRT implantation according to the current guidelines of the American College of Cardiology/American Heart Association/Heart Rhythm Society.1

More recently, baseline PR interval has been invoked as an additional factor that may affect CRT outcomes.2 A prolonged PR interval is a marker of a ventricular substrate that is less amenable to resynchronization. It also reflects a combination of intrinsic intra-atrial and atrioventricular conduction which impacts diastolic filling time.2,3 There are no clear evidence and explanation why PR prolongation might contribute to the outcome of CRT patients. Nonetheless, there is controversial evidence in literature regarding the association between baseline PR prolongation and outcomes of HF patients who require CRT implantation. Some studies implied that PR prolongation was associated with higher morbidity and mortality amongst these patients,2,4-7 while others suggested it is associated with favorable outcomes.8-10 However, a systematic literature review and meta-analysis of the association between PR interval and CRT outcome have not been made.

We have first conducted a systematic literature review and meta-analysis to comprehensively analyze whether baseline
PR prolongation in comparison with normal PR interval is associated with outcomes in CRT-dependent HF patients by assessing all-cause mortality, HF hospitalization rate, and composite outcome as our interest.

Method

Search strategy

Two investigators (NP and TR) independently searched for published studies indexed in MEDLINE and EMBASE databases from inception to January 2017 using a search strategy that included the terms “PR interval” and “cardiac resynchronization therapy” described in online supplementary data. Only English language publications were included. A manual search for additional pertinent studies and review articles using references from retrieved articles was also made.

Inclusion criteria

The eligibility criteria included the following:

1. Cohort study (prospective or retrospective) reporting incident of all-cause mortality, HF hospitalization, or composite outcome, after the CRT and the corresponding index date for controls.
2. Relative risk, hazard ratio, incidence ratio, or standardized incidence ratio with 95% confidence intervals or sufficient raw data for the calculation were provided.
3. Participants without PR prolongation were used as controls.

Study eligibility was independently determined by two investigators (NP and TR) and differences were resolved by mutual consensus. A Newcastle-Ottawa quality assessment scale was used to evaluate each study in three domains: recruitment and selection of the participants, similarity and comparability between the groups, and ascertainment of the outcome of interest among cohort studies.

Data extraction

A standardized data collection form was used to obtain the following information from each study: title of study, name of first author, year of study, year of publication, country of origin, number of participants, demographic data of participants, method used to identify cases and controls, method used to diagnose the outcomes of interest (all-cause mortality, HF hospitalization rate and composite outcome), and average duration of follow-up with confounders that were adjusted and adjusted effect estimates with 95% confidence interval and covariates that were adjusted in the multivariable analysis.

To ensure accuracy, all investigators independently performed this data extraction process. Any data discrepancy was resolved by referring back to the original articles.

Statistical analysis

We performed a meta-analysis of the included cohort studies using a random-effects model. The extracted studies were excluded from the analysis if they did not present an outcome in each intervention group or did not have enough information required for continuous data comparison. We pooled the point estimates from each study using the generic inverse-variance method of DerSimonian and Laird. The heterogeneity of effect size estimates across these studies was quantified using the I² statistic. The I² statistic ranges in value from 0 to 100% (I² < 25%, low heterogeneity; I² = 25%–50%, moderate heterogeneity; and I² > 50%, substantial heterogeneity).

A sensitivity analysis was performed to assess the influence of the individual studies on the overall results by omitting one study at a time. Meta-regression was performed to explore source of heterogeneity. Publication bias was assessed using funnel plot and Egger’s regression test (p < 0.05 was considered significant). All data analyses were performed using the Stata/SE 14.1 software from StataCorp LP.

Results

Description of the included studies

Our search strategy yielded 580 potentially relevant articles (82 articles from EMBASE and 498 articles from MEDLINE). After exclusion of 204 duplicated articles, 376 underwent title and abstract review. Three hundred and seventy articles were excluded at this stage since they were not cohort studies, did not report the outcome of interest (incidence of death/HF hospitalization) or were not conducted in patients with CRT, leaving six for full-length article reviews. Therefore, six retrospective cohort studies with 17,432 normal PR and 4,278 prolonged PR patients were included in this meta-analysis. Figure 1 outlines the search and literature review process. The clinical characteristics and summary of the included studies are described in Table 1.

Quality assessment of the included studies

Newcastle-Ottawa scales of the included studies are described in Table 2. The Newcastle-Ottawa scale uses a star system (0 to 9) to evaluate the included studies on three domains: selection, comparability, and outcomes. Higher scores represent higher study quality. Intra-study risks of bias of the included studies are also described in Table 3.

Meta-analysis results

Six studies2,4,7,8,15,16 from January 1991 to May 2017 were included in this meta-analysis. All-cause mortality rate is available in four studies2,4,7,16 that involved 17,432 normal PR and 4,278 prolonged PR. All four studies revealed an increased death rate among patients with prolonged PR interval but with of the four achieving statistical significance. The pooled analysis demonstrates a statistically significant increased risk of all-cause mortality in patients with prolonged PR interval compared to participants without prolonged PR interval with the pooled risk ratio of 1.34 (95% confidence interval: 1.08–1.67, p < 0.01). The statistical heterogeneity was substantial with I² of 57.0%. Forest plot of this meta-analysis is shown in Figure 2A.
HF hospitalization is available in two studies [2, 4] involving 16,152 normal PR and 3,031 prolonged PR. Both studies achieved statistical significance. HF hospitalization pooled risk ratio is 1.30 (95% confidence interval: 1.16–1.45, \( p < 0.01 \)). The statistical heterogeneity was low with \( I^2 \) of 6.6%. Forest plot of this meta-analysis is shown in Figure 2B.

Composite outcome (all-cause mortality and HF hospitalization) is available in four studies [2, 4, 8, 15] involving 17,001 normal PR and 3,866 prolonged PR. All four studies revealed an increased death rate among patients with prolonged PR interval with two achieving statistical significance. In composite outcome, the pooled analysis also demonstrated a statistically significant increased composite outcome in CRT patients with prolonged PR interval compared to participants without prolonged PR interval with the pooled risk ratio of 1.21 (95% confidence interval: 1.13–1.30, \( p < 0.01 \)). The statistical heterogeneity was low with \( I^2 \) of 0%. Forest plot of this meta-analysis is shown in Figure 2C.

**Sensitivity analysis**

To assess the stability of the results of the meta-analysis, we conducted a sensitivity analysis by excluding one study at a time. None of the results was significantly altered, indicating that our results were robust (supplementary document 2). However, after exclusion of Freidman et al., the heterogeneity decreased from 57.0% to 0% (supplementary document 3).

Given moderate heterogeneity (\( I^2 = 57.0\% \) among all-cause mortality meta-analysis results, meta-regression (supplementary document 3) showed non-significant changes in all-cause mortality in PR interval > 230 msec compared with PR interval > 200 msec with risk ratio of 0.73 (95% confidence interval: 0.43–1.23, \( p = 0.123 \)).

**Publication bias**

To investigate potential publication bias, we examined the funnel plot with pseudo 95% confidence limits of the included studies in assessing change in log risk ratio of death or composite outcome (Figure 3). The vertical axis represents study size (standard error) while the horizontal axis represents effect size (log risk ratio). From this plot, bias is present because there is asymmetrical distribution of studies on both sides of the mean. The Egger's test was significant (\( p < 0.05 \)). However, using the trim and fill methods in the random-effects model, there was no difference of the imputed risk ratio and its 95% confidence interval.

**Discussion**

The evidence provided in this systematic review and meta-analysis shows that a prolonged PR interval is significantly associated with an increased risk for all-cause mortality, composite outcome, and HF hospitalization of patients with CRT.

Prolongation of PR interval, also known as first-degree atrioventricular block, is independently associated with increased risk for mortality and atrial fibrillation in the general population. Even though correlation of PR interval with CRT response was conflicted in previous studies, our meta-analysis confirms the negative effect on clinical outcome in patients with prolonged PR interval. According to the Comparison of Medical Therapy, Pacing and Defibrillation in Heart Failure (COMPANION) trial, around 50% of patients with CRT have prolonged PR interval. In addition, patients with CRT and prolonged PR interval are more likely to have ischemic cardiomyopathy, wider QRS complexes, more
Table 1 – The clinical characteristics and summary of the included studies

<table>
<thead>
<tr>
<th>First author</th>
<th>Friedman</th>
<th>Januszewicz</th>
<th>Kronborg</th>
<th>Olshansky</th>
<th>Lee</th>
<th>Rickard</th>
</tr>
</thead>
<tbody>
<tr>
<td>Country of Origin</td>
<td>USA</td>
<td>USA</td>
<td>Denmark</td>
<td>USA</td>
<td>USA</td>
<td>USA</td>
</tr>
<tr>
<td>Year</td>
<td>2016</td>
<td>2015</td>
<td>2010</td>
<td>2012</td>
<td>2014</td>
<td>2017</td>
</tr>
<tr>
<td>Study type</td>
<td>Retrospective cohort study</td>
<td>Retrospective cohort study</td>
<td>Retrospective cohort study</td>
<td>Retrospective cohort study</td>
<td>Retrospective cohort study</td>
<td>Retrospective cohort study</td>
</tr>
<tr>
<td>Participants description</td>
<td>Patients who underwent CRT (LVEF ≤ 35% and QRS ≥ 120)</td>
<td>Patients who underwent CRT (LVEF ≤ 35%, QRS &gt; 120, NYHA III, IV)</td>
<td>Patients who underwent CRT (LVEF ≤ 35, QRS ≥ 120 and NYHA III, IV)</td>
<td>Patients who underwent CRT (LVEF ≤ 35, QRS ≥ 120 and NYHA III, IV)</td>
<td>Patients who underwent CRT (LVEF ≤ 35, QRS ≥ 120)</td>
<td>Patients who underwent CRT (LVEF ≤ 35, QRS ≥ 120)</td>
</tr>
<tr>
<td>Median duration of follow up (Months)</td>
<td>34</td>
<td>30.1</td>
<td>30</td>
<td>15.95</td>
<td>52.4</td>
<td>61.2</td>
</tr>
<tr>
<td>Definition of prolonged PR</td>
<td>≥ 230 ms</td>
<td>≥ 200 ms</td>
<td>≥ 200 ms</td>
<td>≥ 200 ms</td>
<td>≥ 200 ms</td>
<td>≥ 200 ms</td>
</tr>
<tr>
<td>Number of patients with prolonged PR</td>
<td>2906</td>
<td>125</td>
<td>208</td>
<td>638</td>
<td>204</td>
<td>197</td>
</tr>
<tr>
<td>Number of patients with not prolonged PR</td>
<td>15994</td>
<td>158</td>
<td>232</td>
<td>574</td>
<td>199</td>
<td>275</td>
</tr>
<tr>
<td>Mean age of patients</td>
<td>75.37</td>
<td>66.00</td>
<td>66.00</td>
<td>65.56</td>
<td>66.72</td>
<td>65.10</td>
</tr>
<tr>
<td>confounder adjustment</td>
<td>age, race, QRS, intraventricular conduction, non-ischemic cardiomyopathy, NYHA, HF duration, eGFR, BUN, SBP, sex, RBBB, ischemic cardiomyopathy, AF, medications</td>
<td>age, sex, NYHA, LVEF, LBBB, QRS, HR, SBP, DBP, ischemic status, comorbidities, medication</td>
<td>age, sex, NYHA, LVEF, LBBB, QRS, HR, SBP, DBP, ischemic status, comorbidities, medication</td>
<td>age, sex, NYHA, LVEF, LBBB, QRS, HR, SBP, DBP, ischemic status, comorbidities, medication</td>
<td>age, sex, NYHA, LVEF, LBBB, QRS, HR, SBP, DBP, ischemic status, comorbidities, medication</td>
<td>age, sex, NYHA, LVEF, LBBB, QRS, HR, SBP, DBP, ischemic status, comorbidities, medication</td>
</tr>
</tbody>
</table>

AF: atrial fibrillation; BUN: blood urea nitrogen; HF: heart failure; Cr: creatinine; CRT: cardiac resynchronization therapy; DM: diabetes mellitus; DBP: diastolic blood pressure; eGFR: estimated glomerular filtration rate; HR: heart rate; ICD: implanted cardiac defibrillator; LVEF: left ventricular ejection fraction; MR: mitral regurgitation; NYHA: New York Heart Association; PASP: pulmonary artery systolic pressure; RBBB: right bundle branch block; RV: right ventricular; SBP: systolic blood pressure.

Table 2 – Newcastle–Ottawa scales of the included studies

<table>
<thead>
<tr>
<th>Study</th>
<th>selection</th>
<th>comparability</th>
<th>outcome</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>representativeness</td>
<td>selection of the non-exposed cohort</td>
<td>ascertainment</td>
</tr>
<tr>
<td>Friedman</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>Januszewicz</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>Kronborg</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>Olshansky</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>Ying-Hsiang</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>Rickard</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
</tbody>
</table>

The pathophysiology of PR prolongation causing adverse outcomes is explained by decreased ventricular filling time leading to decreased stroke volume. It can also induce ineffective mitral valve closure, causing diastolic mitral valve regurgitation, which is known to be associated with unfavorable outcomes in left ventricular dysfunction. The study results of Gervais et al. show that after CRT placement, there is a marked subsequent shortening of the mean PR interval, which suggests that CRT cures atrioventricular dyssynchrony. However, our result still shows worse outcome among patients with prolonged PR interval compared to normal PR interval. The reasons for PR interval affecting CRT outcome are uncertain. In general, prolonged PR interval reflects either intrinsic intra-atrial or atrioventricular conduction defect. Thus, CRT may facilitate AV synchrony to mitigate diastolic...
AV valve regurgitation and improve diastolic function.\textsuperscript{19} On the other hand, with the presence of intra-atrial conduction disturbance, CRT implantation could have deleterious impact on these patients as it shortens the appropriate PR interval and causes paradoxical effect, leading to worsening heart failure.\textsuperscript{20} Alternatively, PR prolongation may simply be a rough marker of “sicker” heart failure patients.\textsuperscript{17,21,22}

In current heart failure guidelines, the duration of QRS, the type of bundle branch block and the presence of atrial fibrillation have been utilized as criteria for pacemaker device implantation.\textsuperscript{23} Also, CRT has a range of effects which has promoted interest in refining selection criteria for this important therapy. In our analysis, we imply that the PR interval is a promising prognostic marker in patients with heart failure requiring CRT. Thus, PR interval may also be a valuable adjunctive selection criteria.

As our study has substantial heterogeneity in all cause mortality, we performed sensitivity analysis and found that after exclusion of Friedman et al.,\textsuperscript{2} the heterogeneity decreased from 57.0% to 0%. We concluded that the most likely explanation could be from the definition criteria of the recruited studies. Friedman is the only study that defined prolonged PR as more than 230 msec whereas every other study defined prolonged PR as more than 200 msec. Therefore, a meta-regression was conducted to investigate the statistical significance of PR definition affecting the results. However, meta-regression showed non-significant changes in all-cause mortality in PR interval > 230 msec compared with PR interval > 200 msec.

Our study has some limitations. Despite the fact that our funnel plot does not show biased data set, there are only six studies included in the analysis. In addition, PR prolongation is generally defined as PR interval exceeding 200 milliseconds. However, among the six included studies, there is only one study that defines prolonged PR interval as 230 ms and above.\textsuperscript{2} Given the total number of subjects, the heterogeneity of sample is small. While there are other possible predictor variables that are not included in this study, they were already analyzed in Rickard et al.\textsuperscript{24} Lastly, instead of using cardiac cause-specific mortality, all-cause mortality was used as outcome of interest in the included studies, which might overestimate the total outcome.

### Conclusion

In conclusion, among patients requiring CRT, prolonged PR interval is an independent indicator for all-cause mortality, HF hospitalization, and composite outcome. Our result suggests that PR interval should be considered as one of the important predictors of CRT response when addressing risk stratification.

### Acknowledgement

We would like to thank Elysse Tom, MD for critical reading.

### Author contributions

Conception and design of the research and Statistical analysis: Rattanawong P; Acquisition of data: Prasitlumkum N, Riangwiwat T, Kanjanahattakij N, Chongsathidkiet P; Analysis and interpretation of the data: Rattanawong P, Prasitlumkum N, Riangwiwat T, Kanjanahattakij N, Vutthikraivit W, Chongsathidkiet P; Writing of the manuscript: Prasitlumkum N, Riangwiwat T, Vutthikraivit W; Critical revision of the manuscript for intellectual content: Simpson RJ.

### Potential Conflict of Interest

No potential conflict of interest relevant to this article was reported.

### Sources of Funding

There were no external funding sources for this study.

### Study Association

This study is not associated with any thesis or dissertation work.

### Ethics approval and consent to participate

This article does not contain any studies with human participants or animals performed by any of the authors.
**Figure 2** – Forest plot of the included studies assessing the association between prolonged PR and risk of all-cause mortality (2A), HF hospitalization (2B), and composite outcome (2C).

### A) All-cause mortality

<table>
<thead>
<tr>
<th>Study, Year</th>
<th>Risk Ratio (95% CI)</th>
<th>%Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Freidman et al, 2016</td>
<td>1.12 (1.02, 1.22)</td>
<td>42.35</td>
</tr>
<tr>
<td>Januszkiewicz et al, 2015</td>
<td>1.48 (1.01, 2.17)</td>
<td>19.02</td>
</tr>
<tr>
<td>Kronborg et al, 2010</td>
<td>1.67 (1.14, 2.44)</td>
<td>19.38</td>
</tr>
<tr>
<td>Lee et al, 2014</td>
<td>1.45 (0.99, 2.12)</td>
<td>19.25</td>
</tr>
<tr>
<td>Overall (i-squared = 57.0%, p = 0.072)</td>
<td>1.34 (1.08, 1.67)</td>
<td>100.00</td>
</tr>
</tbody>
</table>

#### Heterogeneity
- Heterogeneity chi-squared = 6.98 (d.f. = 3)
- Estimate of between-study variance Tau-squared = 0.0277
- Test of overall effect: z = 2.62 (p = 0.009)

### B) Heart failure hospitalization

<table>
<thead>
<tr>
<th>Study, Year</th>
<th>Risk Ratio (95% CI)</th>
<th>%Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Freidman et al, 2016</td>
<td>1.28 (1.18, 1.39)</td>
<td>93.32</td>
</tr>
<tr>
<td>Januszkiewicz et al, 2015</td>
<td>1.60 (1.06, 2.43)</td>
<td>6.68</td>
</tr>
<tr>
<td>Overall (i-squared = 6.6%, p = 0.301)</td>
<td>1.30 (1.16, 1.45)</td>
<td>100.00</td>
</tr>
</tbody>
</table>

#### Heterogeneity
- Heterogeneity chi-squared = 1.07 (d.f. = 1)
- Estimate of between-study variance Tau-squared = 0.0016
- Test of overall effect: z = 4.67 (p = 0.000)

### C) Composite outcome

<table>
<thead>
<tr>
<th>Study, Year</th>
<th>Risk Ratio (95% CI)</th>
<th>%Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Freidman et al, 2016</td>
<td>1.23 (1.13, 1.33)</td>
<td>72.90</td>
</tr>
<tr>
<td>Januszkiewicz et al, 2015</td>
<td>1.20 (0.78, 1.85)</td>
<td>2.54</td>
</tr>
<tr>
<td>Olshansky et al, 2012</td>
<td>1.14 (0.97, 1.34)</td>
<td>18.52</td>
</tr>
<tr>
<td>Rickard et al, 2017</td>
<td>1.34 (1.01, 1.77)</td>
<td>6.04</td>
</tr>
<tr>
<td>Overall (i-squared = 0.0%, p = 0.773)</td>
<td>1.21 (1.13, 1.30)</td>
<td>100.00</td>
</tr>
</tbody>
</table>

#### Heterogeneity
- Heterogeneity chi-squared = 1.12 (d.f. = 3)
- Estimate of between-study variance Tau-squared = 0.000
- Test of overall effect: z = 5.53 (p = 0.000)
Figure 3 – Funnel plot of prolonged PR and risk of all-cause mortality (3A), HF hospitalization (3B), and composite outcome (3C). Circles represent the observed published studies.
References


Supplementary Document 1 – Search strategy and keywords

Search strategy and keywords

EMBASE
Searching term:
‘pr interval’ AND ‘cardiac resynchronization therapy’ AND [humans]/lim AND [english]/lim AND [clinical study]/lim

Pubmed
Searching term:
‘pr interval’[All Fields] AND “cardiac resynchronization therapy” [All Fields] NOT “case report”[All Fields]

A) All-cause mortality
Meta-analysis random-effects estimates (exponential form)
Study omitted

Freidman et al, 2016
Januszkwicz et al, 2015
Kronborg et al, 2009
Ying-Hsiang et al, 2014

1.01 1.08 1.34 1.67 1.91

B) Heart failure hospitalization
Meta-analysis random-effects estimates (exponential form)
Study omitted

Freidman et al, 2016
Januszkwicz et al, 2015

1.06 1.16 1.30 1.45 2.43

C) Composite outcome
Meta-analysis random-effects estimates (exponential form)
Study omitted

Freidman et al, 2016
Januszkwicz et al, 2015
Olshansky et al, 2012
Rickard et al, 2017

1.04 1.13 1.21 1.30 1.36

Supplementary Document 2 – Plot of sensitivity analysis of all-cause mortality (S2A), HF hospitalization (S2B), and composite outcome (S2C).
Supplementary Document 3 – Meta-regression of PR definition.